

Multivariate distributions

Question 1: Data and distributional assumptions

(a) Let us assume that we are given data *with all metric columns* of the following form:

| | | | | |
|----------|----------|----------|----------|----------|
| | M_1 | M_2 | \cdots | M_m |
| 1: | x_{11} | x_{12} | \cdots | x_{1n} |
| 2: | x_{21} | x_{22} | \cdots | x_{2n} |
| \vdots | \vdots | \vdots | \ddots | \vdots |
| n: | x_{n1} | x_{n2} | \cdots | x_{nn} |

How would, i.e. as what mathematical objects and using which probabilistic assumptions, would we model the elements of this data to then be able to make inferences about the behaviour/characteristics of new row-wise observations, like $[x_{(n+1)1}, x_{(n+1)2}, \dots, x_{(n+1)m}]$?

(b) Show that the arithmetic mean is an unbiased estimate of the expected value, *given that we are viewing the points we are averaging over as realizations of random variables whose distributions all have the same expected value.*

Do we additionally need to assume that the random variables of which we have realizations are i.i.d.? Explain your answer.

(c) Given the setting of (a), consider the case $m = 1$, i.e. that we only have the data of column M_1 , but are otherwise making the same modelling choices. Show that

$$\hat{\sigma}^2 = \frac{1}{n-1} \sum_{i=1}^n \left(x_i - \frac{1}{n} \sum_{i=1}^n x_i \right)^2$$

is an unbiased estimate for the variance of the distribution we are assuming.

Why are we looking only at the case of $m = 1$ here instead of considering the x_i s to be vectors in \mathbb{R}^m in the above equation?

Question 2: Eigenvalue decomposition

Consider the random vector $\mathbf{x} = (x_1, x_2)^T$ with covariance

$$\Sigma_{\mathbf{x}} = \begin{pmatrix} 2 & 2 \\ 2 & 5 \end{pmatrix}$$

a) Determine the eigenvalues λ_1 and λ_2 and the (normalized) eigenvectors of the matrix $\Sigma_{\mathbf{x}}$.

- b) Use the result of (a) to determine a random vector $\mathbf{y} = (\mathbf{y}_1, \mathbf{y}_2)^T$, whose components \mathbf{y}_1 and \mathbf{y}_2 are linear combinations of \mathbf{x}_1 and \mathbf{x}_2 and vor which additionally holds that

$$\text{Cov}(\mathbf{y}) = \mathbf{\Lambda} = \begin{pmatrix} \lambda_1 & 0 \\ 0 & \lambda_2 \end{pmatrix}.$$

Question 3: Multivariate normal distribution

Let $\mathbf{x} = (\mathbf{x}_1, \dots, \mathbf{x}_p)^T$ be a p -dimensional multivariate-normal distributed random vector. The corresponding density is given by

$$f(\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^p |\mathbf{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \mathbf{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right),$$

where $\boldsymbol{\mu}$ denotes the expected value $\mathbb{E}[\mathbf{x}]$ and $\mathbf{\Sigma}$ the covariance $\text{Cov}(\mathbf{x})$.

- a) Write out the form of this density for the the case $p = 2$, using the parameters $\sigma_i^2 = \text{var}(\mathbf{x}_i)$, $i = 1, 2$, and $\rho = \frac{\text{cov}(\mathbf{x}_1, \mathbf{x}_2)}{\sigma_1 \sigma_2}$. Conclude from this that \mathbf{x}_1 and \mathbf{x}_2 are independent if they are uncorrelated.
- b) Plot the density for $\boldsymbol{\mu} = \mathbf{0}$, $\sigma_1 = 1$, $\sigma_2 = 3$ and different values of ρ using R. (Tip: The function `persp` in combination with the function `manipulate` from the package of the same name is well suited for this).

Question 4: Dermination of marginal distributions

Consider the random variable $\mathbf{z} = (\mathbf{y} \ \mathbf{x})^T \sim \mathcal{N}_{q+p}(\boldsymbol{\mu}, \mathbf{\Sigma})$, with

$$\boldsymbol{\mu} = \begin{pmatrix} \boldsymbol{\mu}_y \\ \boldsymbol{\mu}_x \end{pmatrix}, \quad \mathbf{\Sigma} = \begin{pmatrix} \boldsymbol{\Sigma}_y & \boldsymbol{\Sigma}_{yx} \\ \boldsymbol{\Sigma}_{xy} & \boldsymbol{\Sigma}_x \end{pmatrix}.$$

Derive the marginal distributions for the component vectors \mathbf{y} and \mathbf{x} .